

談雙曲線正割分配之一族

A Note on A Family of Hyperbolic Secant Distributions

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Abstract — We show that a class of hyperbolic secant distributions is not a member of exponential family. A related family of distributions is proved to be complete.

I. Introduction

By a hyperbolic secant distribution we mean a density function of the form

$$f_{\theta}(x) = \frac{\theta}{\pi \cosh x \theta}, \quad -\infty < x < \infty, \quad (1)$$

where θ is a scale parameter such that $0 < \theta < \infty$.

Champernowne [1] had used (1) to calculate the income power (logarithm of the income). The general properties of the above distributions have been reviewed by Lai [2].

The primary aim of this note is to show that although the family of hyperbolic secant distributions is a member of Meixner classes (Meixner [3] & [4]) it does not have an exponential representation. We then show that the family of distributions of $|X_{\theta}|$ is complete (X_{θ} denotes the random variable that corresponds to (1)).

II. Exponential Representation

Theorem 1: The family $\{f_{\theta}(x), -\infty < x < \infty, 0 < \theta < \infty\}$ defined in (1) does not belong to the exponential family of distributions.

Proof: Suppose on the contrary that $\{f_{\theta}(x), 0 < \theta < \infty\}$ belongs to the exponential family, we then could write $f_{\theta}(x)$ in the form

$$f_{\theta}(x) = \frac{\theta}{\pi} \operatorname{sech} x \theta = c(\theta) \exp\{-Q(x)R(\theta)\} h(x), \quad (2)$$

where $Q(x)$, $h(x)$ are functions of x only and $c(\theta)$, $R(\theta)$ are functions of θ only. Now

$$\frac{\pi}{\theta} f_{\theta}(x) = \operatorname{sech} x \theta = \frac{\pi}{\theta} c(\theta) \exp\{-Q(x)R(\theta)\} h(x).$$

Since $\operatorname{sech} x \theta$ is symmetric in x and θ , it is easy to see that we must have $\frac{\pi}{\theta} c(\theta) = h(\theta)$, $Q(x) = R(x)$ and hence

$$\operatorname{sech} x \theta = h(x)h(\theta) \exp\{-Q(x)Q(\theta)\}. \quad (3)$$

Put $\theta = 0$ in (3) we obtain

$$h(x) = \exp \{ Q(x)Q(0) \} / h(0)$$

from which we deduce at once that

$$\operatorname{sech} x \theta = \exp \{ Q(x)Q(0) + Q(\theta)Q(0) - Q(x)Q(\theta) \} / h(0)^2. \quad (4)$$

By putting $\theta = 1$ in (4) we get

$$\operatorname{sech} x = \lambda \exp \beta Q(x) \quad (5)$$

where $\lambda = \exp \{ Q(1)Q(0) \} / h(0)^2 > 0$ and $\beta = Q(0) - Q(1)$.

It is clear that λ and β are constants which are independent of both of x and θ . It follows directly from (5) that

$$Q(x) = -\frac{1}{\beta} \log \lambda \cosh x, \quad Q'(x) = -\frac{1}{\beta} \tanh x. \quad (6)$$

By differentiating both sides of (4) with respect to x we get

$$-\theta \operatorname{sech} x \theta \tanh x \theta = \operatorname{sech} x \theta \{ Q'(x)[Q(0) - Q(\theta)] \}. \quad (7)$$

By letting $x \rightarrow 0$ in (7) we obtain

$$\beta = [Q(0) - Q(\theta)] / \theta^2 \quad (8)$$

Since $Q'(x) = -\frac{1}{\beta} \tanh x$, $Q(x)$ can not be a polynomial of x of degree 2 and hence β cannot be a constant independent of θ . This contradicts our earlier derivation that β is a constant. Hence $\{f_\theta(x): 0 < \theta < \infty\}$ cannot be a one-parameter exponential family of distributions.

We note that since the range of the above distribution does not depend on θ and $\{f_\theta(x): 0 < \theta < \infty\}$ does not belong to the exponential family of distributions, we conclude from the result of Pitman [5] that there exists no sufficient statistic.

III. Completeness

The second aspect of this note is to discuss the problem of completeness of the family $\{g_\theta(y), 0 \leq y < 0, 0 < \theta < \infty\}$, the density functions of $|X_\theta|$ where X_θ , $0 < \theta < \infty$, denote the random variables that correspond to $f_\theta(x)$.

It is well known that a family $F = \{g_\theta(y): \theta \in \Omega\}$ of probability distributions is complete if $E_\theta[\varphi(Y)] = 0$ (for all $g_\theta \in F$) implies $\varphi(y) = 0$ a.e. (see for example, Silvey [6, p29-30]). In our case $g_\theta(y)$ is given by

$$g_\theta(y) = \frac{2\theta}{\pi} \operatorname{sech} \theta y, \quad 0 \leq y < \infty, 0 < \theta < \infty \quad (9)$$

We mention in advance that the discussion presented in this section is somehow heuristic in character.

Theorem 2: The family of distributions defined in (10) indexed by the scale parameter $\theta \in (0, \infty)$ is complete.

Proof: We first note $E_\theta [|\varphi(Y)|] = 0$ implies $E_\theta [|\varphi(Y)|] < \infty$. Therefore for any $\theta > 0$, we have

$$\infty > \int_0^\infty |\varphi(t)| \operatorname{sech} \theta t dt \geq \int_0^R |\varphi(t)| \operatorname{sech} \theta t dt \geq \operatorname{sech} R \theta \int_0^R |\varphi(t)| dt, R > 0, \text{ i.e.}$$

$\varphi(t)$ belongs to L_1 in $(0, R)$ for every $R > 0$.

It is sufficient to show that $\int_0^\infty \varphi(t) \operatorname{sech} \theta t dt = 0$ implies $\varphi(t) = 0$ a.e.

By omitting analytical details we consider

$$\begin{aligned} 2h(\theta) &= \int_0^\infty \varphi(t) \operatorname{sech} \theta t dt, 0 < \theta < \infty \\ &= 2 \int_0^\infty \{ e^{-\theta t} [1 + e^{-2\theta t}]^{-1} \} \varphi(t) dt \\ &= 2 \sum_{n=0}^\infty (-1)^n \int_0^\infty e^{-(2n+1)\theta t} \varphi(t) dt. \end{aligned} \tag{10}$$

Let $\varphi^*(\theta)$ denote the Laplace transform of $\varphi(t)$. Now it is evident

$$h(\theta) = \sum_{n=0}^\infty (-1)^n \varphi^*((2n+1)\theta). \tag{11}$$

Our next step is to find a way to express $\varphi^*(\theta)$ in terms of $h(\theta)$. We shall begin by writing down (11) term-by-term in the manner

$$\begin{aligned} h(\theta) &= \varphi^*(\theta) - \varphi^*(3\theta) + \varphi^*(5\theta) - \varphi^*(7\theta) + \dots \\ h(3\theta) &= \varphi^*(3\theta) - \varphi^*(9\theta) + \varphi^*(15\theta) - \dots \\ h(5\theta) &= \varphi^*(5\theta) - \varphi^*(15\theta) + \dots \\ &\vdots \\ &\dots \dots \dots \text{etc.} \end{aligned}$$

We can verify by means of a careful observation that the above system of equation reduces to

$$\varphi^*(\theta) = \sum_{n=0}^\infty (-1)^n \mu(2n+1) h((2n+1)\theta), \tag{12}$$

where $\mu(\cdot)$ is the Möbius function which can be defined as

$$\mu(n) = \begin{cases} 1 & \text{if } n=1, \\ (-1)^r & \text{if } n \text{ has } r \text{ distinct prime factors,} \\ 0 & \text{otherwise.} \end{cases}$$

For a detailed discussion on Möbius function the reader may consult Van Der Waerden [7, p.114].

Equation (12) shows that $\varphi^*(\theta) = 0$ if $h(\theta) = 0$, i.e. if $E_{\theta}[\varphi(Y)] = \frac{4\theta}{\pi}h(\theta) = 0$.

As we have established earlier, $\varphi(t)$ belongs to L , in $(0, R)$ for every $R > 0$ if $E_{\theta}[\varphi(Y)] = 0$.

Upon applying Lerch Theorem [8, p.62] we deduce at once that $\varphi(t) = 0$ a.e., that is to say $\{g_{\theta}(y), 0 \leq y < \infty, 0 < \theta < \infty\}$ is complete.

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